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8

Appendix

The proofs to Propositions 1 and 2 are in the first section, while the proofs of all lemmas used are in the second section of this appendix.

8.1 Propositions

Proof. (Proposition 1) Suppose \mathcal{H}_0 is true. Without loss of generality and or the sake of clarity, consider a system of two variables, only one lag of only one of the variables and no constant. Under \mathcal{H}_0 and the knowledge of β all the regressors are stationary. Therefore, the asymptotic distributions would be standard. We will show that using super-consistent $\widehat{\beta}$ yields the same limiting distributions. Consider the first equation of the system:

$$\Delta y_{1t} = \theta_{11}(\widehat{z}_{t-1}) + \theta_{21}(\widehat{z}_{t-1})^2 + \theta_{31}(\widehat{z}_{t-1})^3 + \gamma \Delta y_{1t-1} + \tilde{\epsilon}_{1t}, \quad (8.1)$$

where $z_t = \beta' \mathbf{y}_{t-1}$, $\widehat{z}_t = \widehat{\beta}' \mathbf{y}_{t-1}$ and $\tilde{\epsilon}_{1t} = \epsilon_{1t} + \theta_{11}(\widehat{\beta}_2 - \beta_2)y_{2t-1} + \theta_{21}[2z_{t-1}(\widehat{\beta}_2 - \beta_2)y_{2t-1} + 3(\widehat{\beta}_2 - \beta_2)^2 y_{2t-1}^2] + \theta_{31}[3z_{t-1}^2(\widehat{\beta}_2 - \beta_2)y_{2t-1} - 3z_{t-1}(\widehat{\beta}_2 - \beta_2)^2 y_{2t-1}^2 + 5(\widehat{\beta}_2 - \beta_2)^3 y_{2t-1}^3]$. We will prove the result for this equation, the extension to both equations being straightforward but involving much longer and tedious manipulations.

First, note that

$$\sqrt{T} \begin{pmatrix} \widehat{\gamma}_1 - \gamma_1 \\ \widehat{\theta}_{11} - \theta_{11} \\ \widehat{\theta}_{21} - \theta_{21} \\ \widehat{\theta}_{31} - \theta_{31} \end{pmatrix} = T \begin{pmatrix} \sum_{t=1}^T \Delta y_{1t-1}^2 & \sum_{t=1}^T \Delta y_{1t-1} \widehat{z}_{t-1} & \sum_{t=1}^T \Delta y_{1t-1} \widehat{z}_{t-1}^2 & \sum_{t=1}^T \Delta y_{1t-1} \widehat{z}_{t-1}^3 \\ \sum_{t=1}^T \Delta y_{1t-1} \widehat{z}_{t-1} & \sum_{t=1}^T \widehat{z}_{t-1}^2 & \sum_{t=1}^T \widehat{z}_{t-1}^3 & \sum_{t=1}^T \widehat{z}_{t-1}^4 \\ \sum_{t=1}^T \Delta y_{1t-1} \widehat{z}_{t-1}^2 & \sum_{t=1}^T \widehat{z}_{t-1}^3 & \sum_{t=1}^T \widehat{z}_{t-1}^4 & \sum_{t=1}^T \widehat{z}_{t-1}^5 \\ \sum_{t=1}^T \Delta y_{1t-1} \widehat{z}_{t-1}^3 & \sum_{t=1}^T \widehat{z}_{t-1}^4 & \sum_{t=1}^T \widehat{z}_{t-1}^5 & \sum_{t=1}^T \widehat{z}_{t-1}^6 \end{pmatrix}^{-1} \times \frac{1}{\sqrt{T}} \begin{pmatrix} \sum_{t=1}^T \Delta y_{1t-1} \widetilde{\epsilon}_t \\ \sum_{t=1}^T \widehat{z}_{t-1} \widetilde{\epsilon}_t \\ \sum_{t=1}^T \widehat{z}_{t-1}^2 \widetilde{\epsilon}_t \\ \sum_{t=1}^T \widehat{z}_{t-1}^3 \widetilde{\epsilon}_t \end{pmatrix}.$$

Hence, to establish our result, it suffices to show that:

- (a) $\text{plim } T^{-1} \sum_{t=1}^T \Delta y_{1t-1} \widehat{z}_{t-1}^l = \text{plim } T^{-1} \sum_{t=1}^T \Delta y_{1t-1} z_{t-1}^l, \forall l = 1, 2, 3.$
- (b) $\text{plim } T^{-1} \sum_{t=1}^T \widehat{z}_{t-1}^k = \text{plim } T^{-1} \sum_{t=1}^T z_{t-1}^k, \forall k = 2, 3, 4, 5, 6.$
- (c) $T^{-1/2} \sum_{t=1}^T \Delta y_{1t-1} \widetilde{\epsilon}_t$ has the same asymptotic distribution of $T^{-1/2} \sum_{t=1}^T \Delta y_{1t-1} \epsilon_t.$
- (d) $T^{-1/2} \sum_{t=1}^T \widehat{z}_{t-1}^h \widetilde{\epsilon}_t$ has the same asymptotic distribution of $T^{-1/2} \sum_{t=1}^T z_{t-1}^h \epsilon_t$ for $h=1,2,3.$

(a) and (b) follow directly from Lemma 3.

To prove (d), note that the expression $\widehat{z}_{t-1}^2 \widetilde{\epsilon}_t$ equals $\widehat{z}_{t-1}^2 \epsilon_{1t}$ plus a number of terms in the form

$$(\widehat{\beta}_2 - \beta_2)^i \sum_{t=1}^T z_{t-1}^k y_{2t-1}^i.$$

Following what was shown in Lemma 3, as long as $i \geq 1$ and $k > 0$, the expression is $o_p(1)$, i.e., the limit when $T \rightarrow \infty$ is zero. If $k = 0$, we need $i \geq 2$. In this case, the inequality is respected in all expressions.

Hence, $T^{-1/2} \sum_{t=1}^T \widehat{z}_{t-1}^h \widetilde{\epsilon}_{1t} = T^{-1/2} \sum_{t=1}^T \widehat{z}_{t-1}^h \epsilon_{1t} + o_p(1)$. Again from Lemma 3, $T^{-1/2} \sum_{t=1}^T \widehat{z}_{t-1}^h \epsilon_{1t} = T^{-1/2} \sum_{t=1}^T z_{t-1}^h \epsilon_{1t} + o_p(1)$, and from here the result follows.

Proof of claim (c) is analogous to the proof to claim (d).

Now, suppose \mathcal{H}_A is true. To prove the consistency we will show that, under \mathcal{H}_A , the F-statistic diverges to infinity. Under the alternative, Δy_{1t} follows Equation (8.1) except for the error, which becomes $\widetilde{\epsilon}_{1t}^* = \widetilde{\epsilon}_{1t} + \frac{1}{6} f^{(4)}(k_t, \psi)(z_{t-1})^4$ for some fixed $k_t \in \mathbb{R}$.

Let $\mathbf{Z}_t = (\widehat{z}_{t-1}, \widehat{z}_{t-1}^2, \widehat{z}_{t-1}^3, \Delta y_{1t-1})'$ and $\mathbf{Z}_T = (\mathbf{Z}_1, \dots, \mathbf{Z}_T)'$. Then, $\text{plim} \left(\frac{1}{T} \mathbf{Z}_T \mathbf{Z}_T' \right)^{-1} = \mathbf{\Omega}$ is unchanged whether \mathcal{H}_0 is true or not. We will show that under \mathcal{H}_A , $T^{-1/2}(\widehat{\theta}_{21} - 0)$ diverges. Let $\widetilde{\mathbf{\Omega}}$ be the relevant partition of $\mathbf{\Omega}$.

Note that

$$\begin{aligned} T^{-1/2} \widehat{\theta}_{21} &= \widetilde{\mathbf{\Omega}} T^{-1/2} \sum_{t=1}^T \widehat{z}_{t-1}^2 \Delta y_{1t} \\ &= T^{1/2} \theta_{21} + \widetilde{\mathbf{\Omega}} T^{-1/2} \sum_{t=1}^T \widehat{z}_{t-1}^2 \widetilde{\epsilon}_{1t}^* \\ &= T^{1/2} \theta_{21} + \widetilde{\mathbf{\Omega}} T^{-1/2} \sum_{t=1}^T \left[\widehat{z}_{t-1}^2 \widetilde{\epsilon}_{1t} + \widehat{z}_{t-1}^2 \frac{1}{6} f^{(4)}(k_t, \psi)(z_{t-1})^4 \right]. \end{aligned}$$

We know, from Assumptions 1 and 4, that $\frac{1}{6} f^{(4)}(k_t, \psi)$ is bounded. So, we can write

$$-K \sum_{t=1}^T \widehat{z}_{t-1}^2 (z_{t-1})^4 < \sum_{t=1}^T \widehat{z}_{t-1}^2 \frac{1}{6} f^{(4)}(k_1, \psi)(z_{t-1})^4 < K \sum_{t=1}^T \widehat{z}_{t-1}^2 (z_{t-1})^4.$$

Pre-multiplying by $T^{-1/2}$, taking limits, and using the results in Lemma 3, we get

$$O_p(1) < \sum_{t=1}^T \widehat{z}_{t-1}^2 \frac{1}{6} f^{(4)}(k_1, \psi)(z_{t-1})^4 < O_p(1).$$

From (d) we know that $T^{-1/2} \sum_{t=1}^T \widehat{z}_{t-1}^2 \widetilde{\epsilon}_{1t}$ is $O_p(1)$. Therefore, we have two limited terms plus $T^{1/2} \theta_{21}$, which will diverge to ∞ , giving us the result. The same argument applies to the F-test, only with lengthier calculations.

Finally, from (b) it is easy to see that $\text{plim} T^{-1}(\widehat{\epsilon}_{1t}^2)$ exists. ■

Proof. (Proposition 2) Again, for the sake of simplicity, let us consider only one lag of only one variable. In addition, without loss of generality, we will assume ψ is scalar. The NLLS problem is

$$\min \frac{1}{2} T^{-1} \sum_{t=1}^T \left[\Delta y_{1t} - f(\widehat{\beta}' \mathbf{y}_{t-1}, \psi) - \gamma \Delta y_{1t-1} \right]^2.$$

The first order conditions are:

$$T^{-1} \sum_{t=1}^T \widehat{\mathbf{s}}_t(\widehat{\psi}, \widehat{\gamma}) = 0,$$

where

$$\widehat{\mathbf{s}}_t(\widehat{\psi}, \widehat{\gamma}) = \left[\Delta y_{1t} - f(\widehat{\beta}' \mathbf{y}_{t-1}, \widehat{\psi}) - \widehat{\gamma} \Delta y_{1t-1} \right] \begin{bmatrix} f_{\psi}(\widehat{\beta}' \mathbf{y}_{t-1}, \widehat{\psi}) \\ \Delta y_{1t-1} \end{bmatrix}.$$

We will always use the hat to make clear whether the function is calculated with $\widehat{\beta}' \mathbf{y}_{t-1}$ or $\beta' \mathbf{y}_{t-1}$.

We can make a mean-value expansion around (ψ, γ) :

$$\sum_{t=1}^T \widehat{\mathbf{s}}_t(\psi, \gamma) + \sum_{t=1}^T \widehat{\mathbf{H}}_t(\widetilde{\psi}, \widetilde{\gamma}) \begin{pmatrix} \psi - \widehat{\psi} \\ \gamma - \widehat{\gamma} \end{pmatrix} = 0,$$

where

$$\widehat{\mathbf{H}}_t(\psi, \gamma) = \begin{bmatrix} \frac{\partial \widehat{\mathbf{s}}_t(\psi, \gamma)}{\partial \psi} \\ \frac{\partial \widehat{\mathbf{s}}_t(\psi, \gamma)}{\partial \gamma} \end{bmatrix}',$$

$(\widetilde{\psi}, \widetilde{\gamma}) = (\widehat{\psi}, \widehat{\gamma}) + t(\psi, \gamma)$, for some $t \in (0, 1)$.

From Lemma 4

$$\text{plim} T^{-1} \sum_{t=1}^T \widehat{\mathbf{H}}_t(\widetilde{\psi}, \widetilde{\gamma}) = \text{plim} T^{-1} \sum_{t=1}^T \mathbf{H}_t(\psi, \gamma) = \mathbf{H}(\psi, \gamma),$$

for some fixed $\mathbf{H}(\psi, \gamma)$.

Therefore,

$$T^{1/2} \begin{pmatrix} \psi - \widehat{\psi} \\ \gamma - \widehat{\gamma} \end{pmatrix} = [\mathbf{H}_t(\psi, \gamma)]^{-1} (-T^{-1/2}) \sum_{t=1}^T \widehat{\mathbf{s}}_t(\psi, \gamma) + o_p(1).$$

All we have to show now is that

$$(T^{-1/2}) \sum_{t=1}^T \widehat{\mathbf{s}}_t(\psi, \gamma) = (T^{-1/2}) \sum_{t=1}^T \mathbf{s}_t(\psi, \gamma) + o_p(1).$$

It is sufficient to show that

$$(a) \quad T^{-1/2} \sum_{t=1}^T \Delta y_{1t} f_\psi(\widehat{\beta}' \mathbf{y}_{t-1}, \widehat{\psi}) = T^{-1/2} \sum_{t=1}^T \Delta y_{1t} f_\psi(\beta' \mathbf{y}_{t-1}, \psi) + o_p(1)$$

$$(b) \quad T^{-1/2} \sum_{t=1}^T f(\widehat{\beta}' \mathbf{y}_{t-1}, \widehat{\psi}) f_\psi(\widehat{\beta}' \mathbf{y}_{t-1}, \widehat{\psi}) = \\ T^{-1/2} \sum_{t=1}^T f(\beta' \mathbf{y}_{t-1}, \psi) f_\psi(\beta' \mathbf{y}_{t-1}, \psi) + o_p(1)$$

Claims (a) and (b) follow directly from Lemma 3 and Assumption 4.

As to the covariance matrix estimator, the proof is standard. Since we only need to use the Law of Large Numbers, the non-stationarity of the variables does not bring any extra complications Wooldridge (2001) [27].

■

8.2 Lemmas

Lemma 1 *Suppose f is a function which is n times continuously differentiable on the closed interval $[a - r, a + r]$ and $n + 1$ times differentiable on the open interval $(a-r, a+r)$. If there exists a positive real constant M_n such that $|f^{(n+1)}(x)| < M_n, \forall x \in (a - r, a + r)$, then*

$$f(x) = f(a) + f'(a) \frac{(x-a)}{1!} + f''(a) \frac{(x-a)^2}{2!} + \dots + f^{(n)}(b) \frac{(x-a)^n}{n!}$$

for some $b \in (a, x)$.

Proof. See Apostol (1967) [1].

■

From Ibragimov and Phillips (2008) [11]:

Theorem 1 *Let $f : \mathbb{R} \rightarrow \mathbb{R}$ be a twice continuously differentiable function such that f' satisfies the growth condition $|f'(x)| \leq K(1 + |x|^\alpha), \forall x \in \mathbf{R}$ for some constants $K > 0$ and $\alpha < 0$. Suppose that u_t and v_t are two linear processes $u_t = \sum_{j=1}^{\infty} \gamma_j \epsilon_{t-j}$ and $v_t = \sum_{j=1}^{\infty} \delta_j \epsilon_{t-j}$ where $\sum_{j=1}^{\infty} j |\gamma_j| < \infty$, $\sum_{j=1}^{\infty} j |\delta_j| < \infty$ and $(\epsilon_t)_{t \in \mathbf{Z}}$ are zero-mean i.i.d. random variables with $E[\epsilon_0^2] < \infty$ and $E[|\epsilon_0|^p] < \infty$ for $p \geq \max(6, 4\alpha)$. Then*

$$\frac{1}{\sqrt{T}} \sum_{t=2}^{[Tr]} f \left(\frac{1}{\sqrt{T}} \sum_{i=1}^{t-1} u_i \right) v_t \xrightarrow{d} \lambda_{uv} \int_0^r f'(\omega_u W(v)) dv + \omega_v \int_0^r f(\omega_u W(v)) d(W(v)),$$

where $\omega_u = E[u_t^2]$, $\omega_v = E[v_t^2]$ and $\lambda_{uv} = \sum_{j=1}^{\infty} E[u_0 v_0]$.

The exact form of the limiting distribution is not relevant for our results. What we need is the following corollary.

Lemma 2 Under the conditions of Theorem 1,

$$T^{-1/2} \sum_{t=2}^T f \left(T^{-1/2} \sum_{i=1}^{t-1} u_i \right) v_t = O_p(1)$$

Note that the derivatives of any polynomial function satisfy the growth condition.

Lemma 3 Let v_t be a stationary process, \mathbf{y}_t be an $I(1)$ cointegrated vector, with cointegration vector $\boldsymbol{\beta}$ and $\widehat{\boldsymbol{\beta}}$ a super-consistent estimate of $\boldsymbol{\beta}$. Let also, for some $d < \infty$, $f : \mathbb{R} \rightarrow \mathbb{R}$ be d times continuously differentiable and the d -th derivative of f be limited. Then, $T^{-1/2} \sum_{t=1}^T f(\boldsymbol{\beta}' \mathbf{y}_t) v_t = T^{-1/2} \sum_{t=1}^T f(\widehat{\boldsymbol{\beta}}' \mathbf{y}_t) v_t + o_p(1)$, and $\text{plim } T^{-1} \sum_{t=1}^T f(\boldsymbol{\beta}' \mathbf{y}_t) = \text{plim } T^{-1} \sum_{t=1}^T f(\widehat{\boldsymbol{\beta}}' \mathbf{y}_t)$.

Proof. For the first result, consider first a two dimensional case $\boldsymbol{\beta}' \mathbf{y}_t = y_{1t} + \beta_2 y_{2t}$. Using Lemma 1 to expand f around $\beta_2 y_{2t}$,

$$\begin{aligned} f(\widehat{\boldsymbol{\beta}}' \mathbf{y}_t) &= f(\boldsymbol{\beta}' \mathbf{y}_t) + f'(\boldsymbol{\beta}' \mathbf{y}_t) (\widehat{\beta}_2 - \beta_2) y_{2t} + \dots + \frac{f^{(d-1)}(\boldsymbol{\beta}' \mathbf{y}_t) (\widehat{\beta}_2 - \beta_2)^{d-1} y_{2t}^{d-1}}{(d-1)!} \\ &\quad + \frac{f^d(\widetilde{\boldsymbol{\beta}}' \mathbf{y}_t) (\widehat{\beta}_2 - \beta_2)^d y_{2t}^d}{d!} \end{aligned}$$

for some $\widetilde{\boldsymbol{\beta}}' \mathbf{y}_t \in (\widehat{\boldsymbol{\beta}}' \mathbf{y}_t, \boldsymbol{\beta}' \mathbf{y}_t)$. Taking the k -th term, such that $3 \leq k \leq d-1$, we have, by Lemma 2,

$$\begin{aligned} (\widehat{\beta}_2 - \beta_2)^k \sum_{t=1}^T f^k(\boldsymbol{\beta}' \mathbf{y}_t) v_t y_{2t}^k / k! &= T^{-(k-1)/2} \left[T(\widehat{\beta}_2 - \beta_2) \right]^k \\ &\quad \times \left[\frac{T^{-(k+1)/2} \sum_{t=1}^T f^k(\boldsymbol{\beta}' \mathbf{y}_t) v_t y_{2t}^k}{k!} \right] \\ &= T^{-(k-1)/2} O_p(1) O_p(1) = o_p(1). \end{aligned}$$

For $k = 2$ we get a $O_p(1)$, but it will be further divided by $T^{1/2}$, giving us an $o_p(1)$. Since the d -th derivative is limited, for some $M \in \mathbb{R}$, the sum of the d -th term is bounded by

$$\begin{aligned} \pm M \sum_{t=1}^T (\widehat{\beta}_2 - \beta_2)^d v_t y_{2t}^d &= \pm M T^{-(d-1)/2} \left[T(\widehat{\beta}_2 - \beta_2) \right]^d \left[T^{-(d+1)/2} \sum_{t=1}^T v_t y_{2t}^d \right] \\ &= T^{-(d-1)/2} O_p(1) O_p(1). \end{aligned}$$

Again, if $d \neq 1$ we have an $o_p(1)$ expression, if $d = 1$, we are back to the $k = 2$ case. Therefore, the only remaining term is the first, which gives us the result.

For the multidimensional case, just repeat the reasoning for each dimension of β' .

The second result is proven by the same line of reasoning. The only difference is that in the end of each expression we will have $T^{-(i-2)/2}O_p(1)O_p(1)$, giving us an $o_p(1)$ except for $i = 1, 2$. But since the expression will be divided by T , we will have $o_p(1)$ for every i .

■

Lemma 4 $\text{plim } T^{-1} \sum_{t=1}^T \widehat{\mathbf{H}}_t(\tilde{\psi}, \tilde{\gamma}) = \text{plim } T^{-1} \sum_{t=1}^T \mathbf{H}_t(\psi, \gamma)$.

Proof. $\widehat{\mathbf{H}}_t(\tilde{\psi}, \tilde{\gamma})$ equals

$$\begin{bmatrix} L & \Delta y_{1t-1} f_{\psi}(\tilde{\beta}' \mathbf{y}_{t-1}, \tilde{\psi}) \\ \Delta y_{1t-1} f_{\psi}(\tilde{\beta}' \mathbf{y}_{t-1}, \tilde{\psi}) & \Delta y_{1t-1}^2 \end{bmatrix}$$

Where $L = f_{\psi\psi}(\tilde{\beta}' \mathbf{y}_{t-1}, \tilde{\psi})(-\Delta \mathbf{y}_{1t} + f(\tilde{\beta}' \mathbf{y}_{t-1}, \tilde{\psi}) + \tilde{\gamma} \Delta y_{1t-1}) + f_{\psi}(\tilde{\beta}' \mathbf{y}_{t-1}, \tilde{\psi})^2$. From Lemma 3 and Assumption 4:

- (a) $\text{plim } T^{-1} \sum_{t=1}^T \Delta y_{1t-1} f_{\psi}(\tilde{\beta}' \mathbf{y}_{t-1}, \tilde{\psi}) = \text{plim } T^{-1} \sum_{t=1}^T \Delta y_{1t-1} f_{\psi}(\beta' \mathbf{y}_{t-1}, \tilde{\psi})$
- (b) $\text{plim } T^{-1} \sum_{t=1}^T f_{\psi\psi}(\tilde{\beta}' \mathbf{y}_{t-1}, \tilde{\psi}) \left[-\Delta y_{1t} + f(\tilde{\beta}' \mathbf{y}_{t-1}, \tilde{\psi}) + \tilde{\gamma} \Delta y_{1t-1} \right] =$
 $\text{plim } T^{-1} \sum_{t=1}^T f_{\psi\psi}(\beta' \mathbf{y}_{t-1}, \tilde{\psi}) \left[-\Delta y_{1t} + f(\beta' \mathbf{y}_{t-1}, \tilde{\psi}) + \tilde{\gamma} \Delta y_{1t-1} \right]$
- (c) $\text{plim } T^{-1} \sum_{t=1}^T f_{\psi}(\tilde{\beta}' \mathbf{y}_{t-1}, \tilde{\psi})^2 = \text{plim } T^{-1} \sum_{t=1}^T f_{\psi}(\beta' \mathbf{y}_{t-1}, \tilde{\psi})^2$

Therefore, we have established that $\text{plim } T^{-1} \sum_{t=1}^T \widehat{\mathbf{H}}_t(\tilde{\psi}, \tilde{\gamma}) = \text{plim } T^{-1} \sum_{t=1}^T \mathbf{H}_t(\tilde{\psi}, \tilde{\gamma})$. Usual nonlinear least squares approach, c.f. Wooldridge (2001) [27], may easily be used to establish $\text{plim } T^{-1} \sum_{t=1}^T \mathbf{H}_t(\tilde{\psi}, \tilde{\gamma}) = \text{plim } T^{-1} \sum_{t=1}^T \mathbf{H}_t(\psi, \gamma)$. These two equalities give us the result.

■